

# The BDe score

A brief derivation

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The parent node set of node  $X_n$  in  $\mathcal{G}$ ,  $\pi_n = \pi_n(\mathcal{G})$ , is the set of all nodes from which an edge points to node  $X_n$  in  $\mathcal{G}$ . The conditional dependence of  $X_n$  on  $\pi_n$  is described by the conditional probability (density)  $P(X_n|\pi_n, \boldsymbol{\theta}_{n,\pi_n})$ , which lies in a chosen function family, and depends on the parameters  $\boldsymbol{\theta}_{n,\pi_n}$ . Consider a set of experimental conditions  $t, 1 \leq t \leq M$ . Let  $X_n(t)$  represent the realization of  $X_n$  in experimental condition  $t$ . We are given a data set  $\mathcal{D}$ , where  $\mathcal{D}_{n,t}$  and  $\mathcal{D}_{(\pi_n,t)}$  are the  $t$ th realizations  $X_n(t)$  and  $\pi_n(t)$  of  $X_n$  and  $\pi_n$ , respectively.

$$P(\mathcal{D}|\mathcal{G}) = \int P(\mathcal{D}|\mathcal{G}, \boldsymbol{\theta})P(\boldsymbol{\theta}|\mathcal{G})d\boldsymbol{\theta}$$

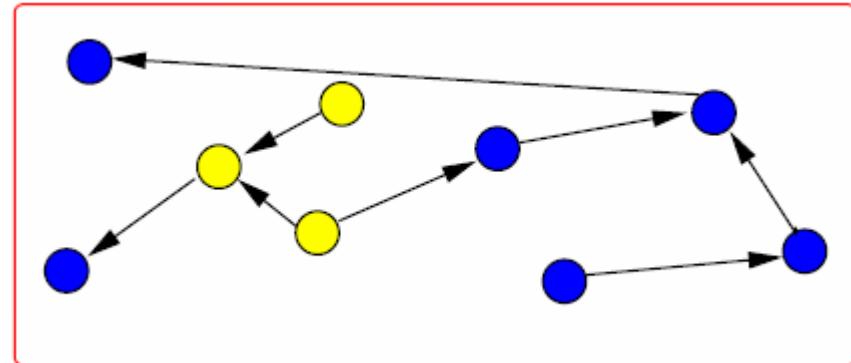
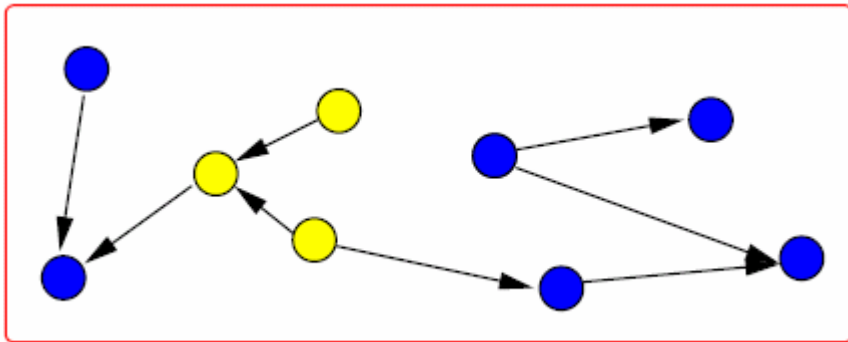
$$P(\mathcal{D}|\mathcal{G}, \boldsymbol{\theta}) = \prod_{n=1}^N \prod_{t=1}^M P\left(X_n(t) = \mathcal{D}_{n,t} | \pi_n(t) = \mathcal{D}_{(\pi_n,t)}, \boldsymbol{\theta}_{n,\pi_n}\right)$$

Parameter modularity:

$$P(\boldsymbol{\theta}_{n,\pi_n}|\mathcal{G}) = P(\boldsymbol{\theta}_{n,\pi_n}|\mathcal{G}') := P(\boldsymbol{\theta}_{n,\pi_n})$$

Parameter independence:

$$P(\boldsymbol{\theta}|\mathcal{G}) = \prod_{n=1}^N P(\boldsymbol{\theta}_{n,\pi_n}|\mathcal{G})$$



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Parameter independence:

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## Decomposition of the marginal likelihood

$$\begin{aligned} P(\mathcal{D}|\mathcal{G}) &= \int P(\mathcal{D}|\mathcal{G}, \boldsymbol{\theta})P(\boldsymbol{\theta}|\mathcal{G})d\boldsymbol{\theta} \\ &= \int \prod_{n=1}^N \prod_{t=1}^M P\left(X_n(t) = \mathcal{D}_{n,t} | \pi_n(t) = \mathcal{D}_{(\pi_n,t)}, \boldsymbol{\theta}_{n,\pi_n}\right) P(\boldsymbol{\theta}_{n,\pi_n}) d\boldsymbol{\theta}_{n,\pi_n} \end{aligned}$$

$$P(\mathcal{D}|\mathcal{G}) = \int P(\mathcal{D}|\mathcal{G}, \boldsymbol{\theta})P(\boldsymbol{\theta}|\mathcal{G})d\boldsymbol{\theta} = \prod_{n=1}^N \Psi(\mathcal{D}_n^{\pi_n})$$

$$\Psi(\mathcal{D}_n^{\pi_n}) = \int \prod_{t=1}^M P\left(X_n(t) = \mathcal{D}_{n,t} | \pi_n(t) = \mathcal{D}_{(\pi_n,t)}, \boldsymbol{\theta}_{n,\pi_n}\right) P(\boldsymbol{\theta}_{n,\pi_n}) d\boldsymbol{\theta}_{n,\pi_n}$$

where  $\mathcal{D}_n^{\pi_n} := \{(\mathcal{D}_{n,t}, \mathcal{D}_{\pi_n,t}) : 1 \leq t \leq m\}$

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Analytically tractable?

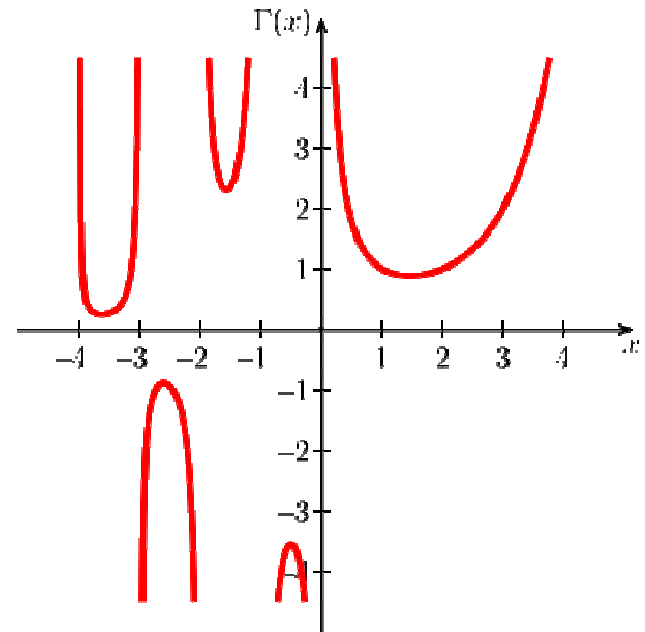
## Mathematical Primer

$$\Gamma(z) = \int_0^{\infty} t^{z-1} \exp(-t) dt$$

$$\Gamma(z + 1) = z\Gamma(z)$$

$$\Gamma(1) = 1$$

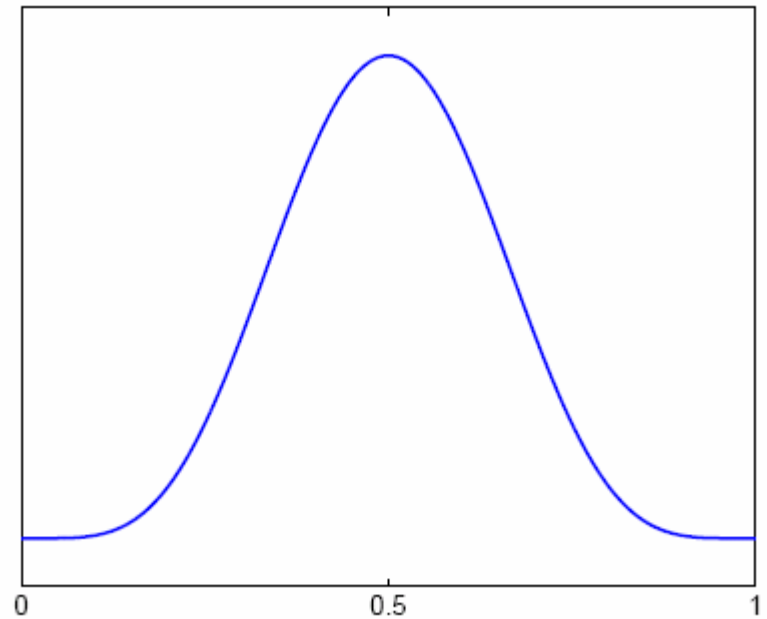
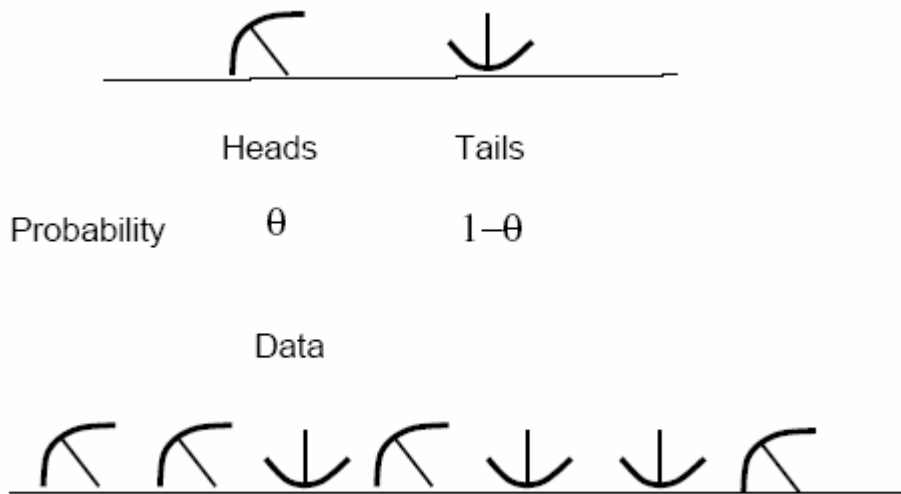
$$\Gamma(n + 1) = n!$$



$$\int_0^1 \theta^{\alpha-1} (1-\theta)^{\beta-1} d\theta = \mathcal{B}(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$$

# Binomial distribution

$$L(\theta) = P(\mathcal{D}|\theta)$$



Likelihood:

$$L(\theta) = P(\mathcal{D}|\theta) = \binom{M}{m} \theta^m (1 - \theta)^{M-m}$$

Total number of observations



Number of heads = sufficient statistics

## Binomial distribution

Tossing a coin,  $\theta \in [0, 1]$ ,  $M$  observations,  $m$  times heads.

Likelihood:

$$L(\theta) = P(\mathcal{D}|\theta) = \binom{M}{m} \theta^m (1 - \theta)^{M-m}$$

Log Likelihood:

$$\log L(\theta) = \log P(\mathcal{D}|\theta) = m \log \theta + (M - m) \log(1 - \theta) + C$$

Maximum Likelihood:

$$\begin{aligned} \frac{d \log L(\theta)}{d\theta} &= \frac{m}{\theta} - \frac{M - m}{1 - \theta} = 0 \\ m(1 - \theta) &= (M - m)\theta \\ \theta &= \frac{m}{M} \end{aligned}$$

Tossing a coin,  $\theta \in [0, 1]$ ,  $M$  observations,  $m$  times heads.

Likelihood:

$$L(\theta) = P(\mathcal{D}|\theta) = \binom{M}{m} \theta^m (1 - \theta)^{M-m}$$

Conjugate Prior:

$$P(\theta) = \frac{1}{Z} \theta^{\alpha-1} (1 - \theta)^{\beta-1}$$

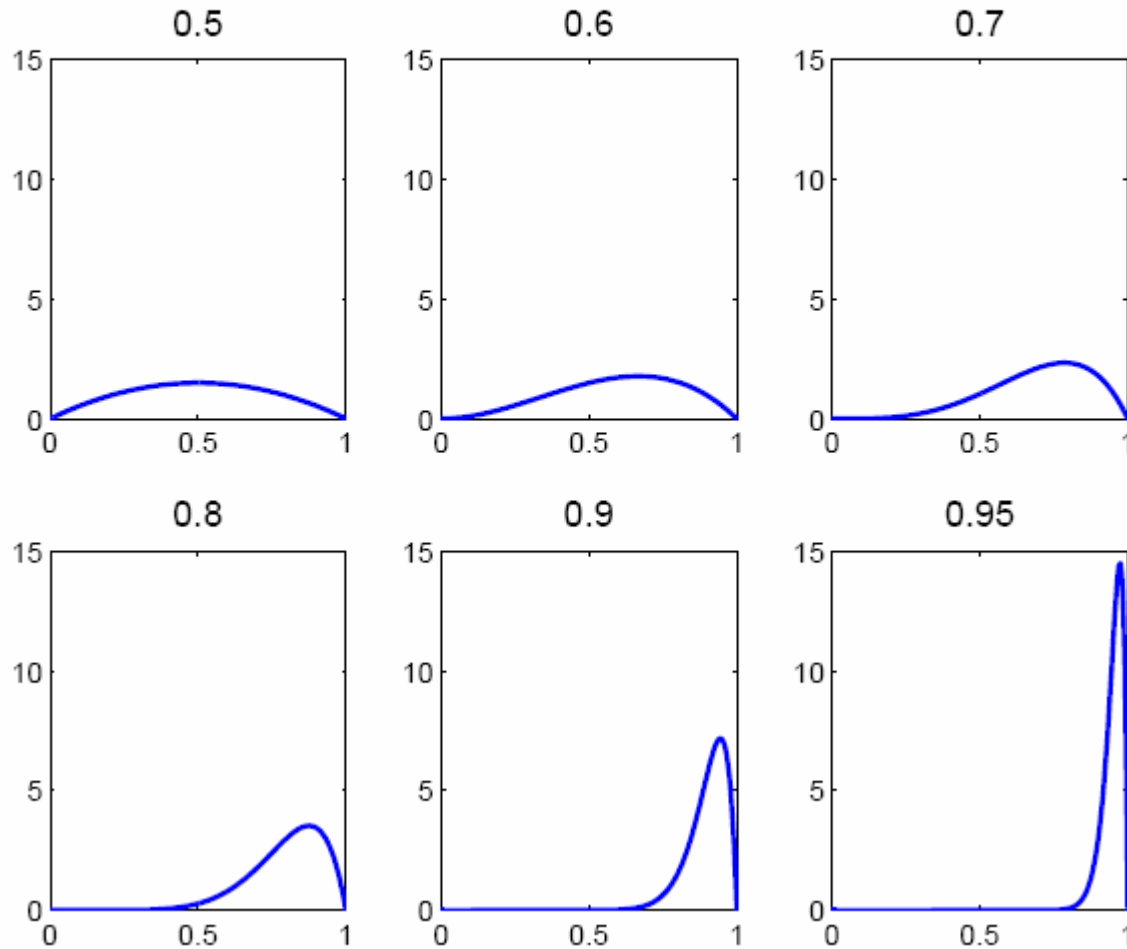
Posterior:

$$\begin{aligned} P(\theta|\mathcal{D}) &\propto P(\mathcal{D}|\theta)P(\theta) \\ &\propto \theta^m (1 - \theta)^{M-m} \times \theta^{\alpha-1} (1 - \theta)^{\beta-1} \\ &= \theta^{(m+\alpha)-1} (1 - \theta)^{(\beta+M-m)-1} \\ &= \theta^{\tilde{\alpha}-1} (1 - \theta)^{\tilde{\beta}-1} \end{aligned}$$

$$\tilde{\alpha} = \alpha + m, \quad \tilde{\beta} = \beta + M - m$$

Beta Prior,  $\beta = 2$ ,  $\mu = \alpha / (\alpha + \beta)$

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$\alpha = \beta = 1 \rightarrow$  uniform prior over  $[0, 1]$

Conjugate Prior:

$$P(\theta) = \frac{1}{Z} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$

Normalization constant:

$$Z = \int_0^1 \theta^{\alpha-1} (1-\theta)^{\beta-1} d\theta = \mathcal{B}(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$$

$$\Gamma(z) = \int_0^{\infty} t^{z-1} \exp(-t) dt$$

$$\Gamma(z + 1) = z\Gamma(z)$$

$$\Gamma(1) = 1$$

$$\Gamma(n + 1) = n!$$

Prior:

$$P(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)}\theta^{\alpha-1}(1 - \theta)^{\beta-1}$$

Posterior:

$$P(\theta|\mathcal{D}) = \frac{\Gamma(\tilde{\alpha} + \tilde{\beta})}{\Gamma(\tilde{\alpha})\Gamma(\tilde{\beta})}\theta^{\tilde{\alpha}-1}(1 - \theta)^{\tilde{\beta}-1}$$

$$\tilde{\alpha} = \alpha + m, \quad \tilde{\beta} = \beta + M - m$$

Posterior:

$$P(\theta|\mathcal{D}) = \frac{\Gamma(\tilde{\alpha} + \tilde{\beta})}{\Gamma(\tilde{\alpha})\Gamma(\tilde{\beta})} \theta^{\tilde{\alpha}-1} (1 - \theta)^{\tilde{\beta}-1}$$

Posterior mean:

$$\begin{aligned} \langle \theta \rangle &= \int_0^1 \theta P(\theta|\mathcal{D}) d\theta \\ &= \frac{\Gamma(\tilde{\alpha} + \tilde{\beta})}{\Gamma(\tilde{\alpha})\Gamma(\tilde{\beta})} \int_0^1 \theta^{\tilde{\alpha}+1-1} (1 - \theta)^{\tilde{\beta}-1} d\theta \end{aligned}$$

Posterior:

$$P(\theta|\mathcal{D}) = \frac{\Gamma(\tilde{\alpha} + \tilde{\beta})}{\Gamma(\tilde{\alpha})\Gamma(\tilde{\beta})} \theta^{\tilde{\alpha}-1} (1 - \theta)^{\tilde{\beta}-1}$$

**Recall:**  $\int_0^1 \theta^{\alpha-1} (1 - \theta)^{\beta-1} d\theta = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$   $\Gamma(z + 1) = z\Gamma(z)$

Posterior mean:

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Posterior:

$$P(\theta|\mathcal{D}) = \frac{\Gamma(\tilde{\alpha} + \tilde{\beta})}{\Gamma(\tilde{\alpha})\Gamma(\tilde{\beta})} \theta^{\tilde{\alpha}-1} (1 - \theta)^{\tilde{\beta}-1}$$

Recall:  $\int_0^1 \theta^{\alpha-1} (1 - \theta)^{\beta-1} d\theta = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$   $\Gamma(z + 1) = z\Gamma(z)$

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## Binomial distribution

Tossing a coin,  $\theta \in [0, 1]$ ,  $M$  observations,  $m$  times heads.

Likelihood:

$$L(\theta) = P(\mathcal{D}|\theta) = \binom{M}{m} \theta^m (1 - \theta)^{M-m}$$

## Multinomial distribution

$K$  outcomes, parameter vector  $\boldsymbol{\theta} = (\theta_1, \dots, \theta_k, \dots, \theta_K) \in \mathcal{S}$ , i.e.  $\theta_k \geq 0 \forall k$ ,  $\sum_{k=1}^K \theta_k = 1$ .  $M$  observations,  $M_k$  in category  $k$ ,  $k \in \{1, \dots, K\}$ .

Likelihood:

$$L(\boldsymbol{\theta}) = P(\mathcal{D}|\boldsymbol{\theta}) = \frac{M!}{\prod_k M_k!} \prod_{k=1}^K \theta^{M_k}$$

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Conjugate Prior:

$$P(\boldsymbol{\theta}) = Dir(\boldsymbol{\theta}) = \frac{\Gamma(\sum_k \alpha_k)}{\prod_k \Gamma(\alpha_k)} \prod_{k=1}^K \theta_k^{\alpha_k - 1}$$

Note:

$$\int_{\boldsymbol{\theta} \in \mathcal{S}} \prod_{k=1}^K \theta_k^{\alpha_k - 1} d\boldsymbol{\theta} = \left( \frac{\Gamma(\sum_k \alpha_k)}{\prod_k \Gamma(\alpha_k)} \right)^{-1}$$

Posterior:

$$P(\boldsymbol{\theta}|\mathcal{D}) \propto P(\mathcal{D}|\boldsymbol{\theta})P(\boldsymbol{\theta}) = \prod_{k=1}^K \theta^{M_k} \prod_{k=1}^K \theta_k^{\alpha_k-1} = \prod_{k=1}^K \theta_k^{M_k+\alpha_k-1} = \prod_{k=1}^K \theta_k^{\tilde{\alpha}_k-1}$$

Marginal likelihood:

$$P(\mathcal{D}|\boldsymbol{\alpha}) = \int_{\mathcal{S}} P(\mathcal{D}|\boldsymbol{\theta})P(\boldsymbol{\theta}|\boldsymbol{\alpha})d\boldsymbol{\theta} = \int_{\mathcal{S}} \prod_{k=1}^K \theta_k^{\tilde{\alpha}_k-1} d\boldsymbol{\theta} = \frac{\prod_k \Gamma(\tilde{\alpha}_k)}{\Gamma(\sum_k \tilde{\alpha}_k)}$$

Posterior:

$$P(\boldsymbol{\theta}|\mathcal{D}) = \frac{P(\mathcal{D}|\boldsymbol{\theta})P(\boldsymbol{\theta}|\boldsymbol{\alpha})}{P(\mathcal{D}|\boldsymbol{\alpha})} = \frac{\Gamma(\sum_k \tilde{\alpha}_k)}{\prod_k \Gamma(\tilde{\alpha}_k)} \prod_{k=1}^K \theta_k^{\tilde{\alpha}_k-1}$$

$$P(\boldsymbol{\theta}|\mathcal{D}) = \frac{\Gamma(\sum_k \tilde{\alpha}_k)}{\prod_k \Gamma(\tilde{\alpha}_k)} \prod_{k=1}^K \theta_k^{\tilde{\alpha}_k - 1} \quad \int_{\boldsymbol{\theta} \in \mathcal{S}} \prod_{k=1}^K \theta_k^{\alpha_k - 1} d\boldsymbol{\theta} = \left( \frac{\prod_k \Gamma(\alpha_k)}{\Gamma(\sum_k \alpha_k)} \right)$$

$$\Gamma(z+1) = z\Gamma(z)$$

Posterior mean:

$$\begin{aligned} \langle \theta_1 \rangle &= \int_{\boldsymbol{\theta} \in \mathcal{S}} \theta_1 P(\boldsymbol{\theta}|\mathcal{D}) d\boldsymbol{\theta} \\ &= \frac{\Gamma(\sum_{k=1}^K \tilde{\alpha}_k)}{\prod_{k=1}^K \Gamma(\tilde{\alpha}_k)} \int_{\boldsymbol{\theta} \in \mathcal{S}} \theta_1^{\tilde{\alpha}_1 + 1 - 1} \prod_{k=2}^K \theta_k^{\tilde{\alpha}_k - 1} d\boldsymbol{\theta} \\ &= \frac{\Gamma(\sum_{k=1}^K \tilde{\alpha}_k)}{\prod_{k=1}^K \Gamma(\tilde{\alpha}_k)} \times \frac{\Gamma(\tilde{\alpha}_1 + 1) \prod_{k=2}^K \Gamma(\tilde{\alpha}_k)}{\Gamma(\tilde{\alpha}_1 + 1 + \sum_{k=2}^K \tilde{\alpha}_k)} \end{aligned}$$

$$P(\boldsymbol{\theta}|\mathcal{D}) = \frac{\Gamma(\sum_k \tilde{\alpha}_k)}{\prod_k \Gamma(\tilde{\alpha}_k)} \prod_{k=1}^K \theta_k^{\tilde{\alpha}_k - 1} \quad \int_{\boldsymbol{\theta} \in \mathcal{S}} \prod_{k=1}^K \theta_k^{\alpha_k - 1} d\boldsymbol{\theta} = \left( \frac{\prod_k \Gamma(\alpha_k)}{\Gamma(\sum_k \alpha_k)} \right)$$

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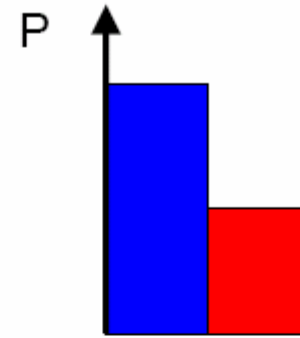
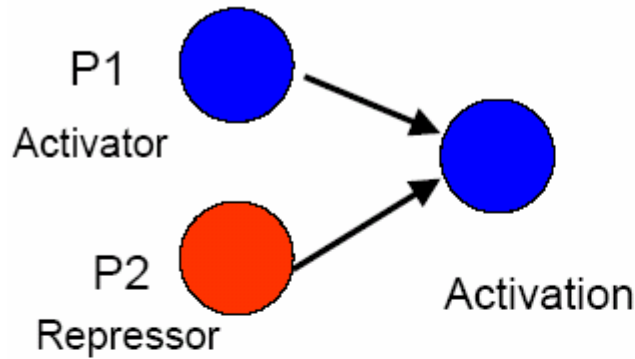
$$P(\boldsymbol{\theta}|\mathcal{D}) = \frac{\Gamma(\sum_k \tilde{\alpha}_k)}{\prod_k \Gamma(\tilde{\alpha}_k)} \prod_{k=1}^K \theta_k^{\tilde{\alpha}_k - 1} \quad \int_{\boldsymbol{\theta} \in \mathcal{S}} \prod_{k=1}^K \theta_k^{\alpha_k - 1} d\boldsymbol{\theta} = \left( \frac{\prod_k \Gamma(\alpha_k)}{\Gamma(\sum_k \alpha_k)} \right)$$

$$\Gamma(z+1) = z\Gamma(z)$$

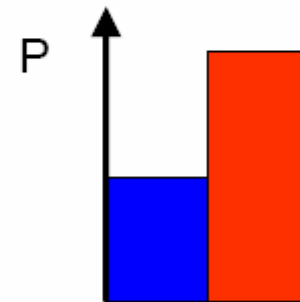
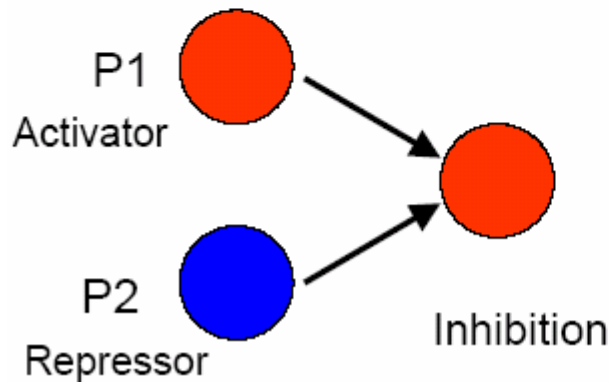
Posterior mean:

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# Discrete Bayesian network



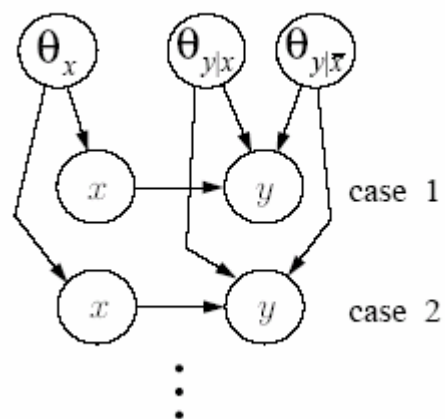
Allow for noise: probabilities



Conditional multinomial distribution

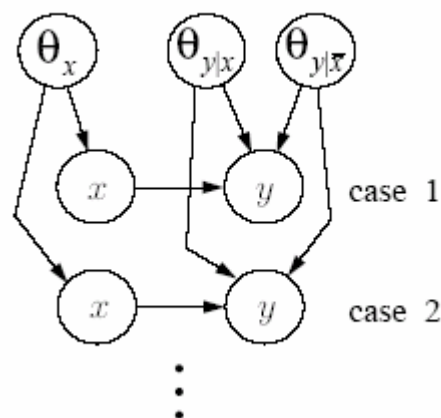
$$P\left(X_n(t) = \mathcal{D}_{n,t} \mid \pi_n(t) = \mathcal{D}_{(\pi_n,t)}, \boldsymbol{\theta}_{n,\pi_n}\right) = \prod_{k=1}^K \theta_{n,\pi_n,k} \mathcal{I}(\mathcal{D}_{n,t} = k)$$

## Discrete Bayesian Networks



$$P\left(X_n(t) = \mathcal{D}_{n,t} \mid \pi_n(t) = \mathcal{D}_{(\pi_n,t)}, \boldsymbol{\theta}_{n,\pi_n}\right) = \prod_{k=1}^K \theta_{n,\pi_n,k} \mathcal{I}(\mathcal{D}_{n,t} = k)$$

## Discrete Bayesian Networks



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$$\begin{aligned} P(\mathcal{D} \mid \mathcal{G}, \boldsymbol{\theta}) &= \prod_{n=1}^N \prod_{t=1}^M P\left(X_n(t) = \mathcal{D}_{n,t} \mid \pi_n(t) = \mathcal{D}_{(\pi_n,t)}, \boldsymbol{\theta}_{n,\pi_n}\right) \\ &= \prod_{n=1}^N \prod_{t=1}^M \prod_{k=1}^K \theta_{n,\pi_n,k} \mathcal{I}(\mathcal{D}_{n,t} = k) \\ &= \prod_n \prod_{\pi_n} \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k}} \end{aligned}$$

Parameter modularity:

$$P(\boldsymbol{\theta}_{n,\pi_n}|\mathcal{G}) = P(\boldsymbol{\theta}_{n,\pi_n}|\mathcal{G}') := P(\boldsymbol{\theta}_{n,\pi_n})$$

Parameter independence:

$$P(\boldsymbol{\theta}|\mathcal{G}) = \prod_{n=1}^N P(\boldsymbol{\theta}_{n,\pi_n}|\mathcal{G})$$

Conjugacy:

$$P(\boldsymbol{\theta}_{n,\pi_n}) = Dir(\boldsymbol{\theta}_{n,\pi_n}|\boldsymbol{\alpha}_{n,\pi_n}) = \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\prod_k \Gamma(\alpha_{n,\pi_n,k})} \prod_k [\theta_{n,\pi_n,k}]^{\alpha_{n,\pi_n,k}-1}$$

Likelihood:

$$P(\mathcal{D}|\mathcal{G}, \boldsymbol{\theta}) = \prod_n \prod_{\pi_n} \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k}}$$

Recall:

$$Dir(\boldsymbol{\theta}) = \frac{\Gamma(\sum_k \alpha_k)}{\prod_k \Gamma(\alpha_k)} \prod_{k=1}^K \theta_k^{\alpha_k-1}$$

Marginal likelihood:

$$\begin{aligned} P(\mathcal{D}|\mathcal{G}) &= \int P(\mathcal{D}|\boldsymbol{\theta}, \mathcal{G}) P(\boldsymbol{\theta}|\mathcal{G}) d\boldsymbol{\theta} \\ &= \prod_n \prod_{\pi_n} \int \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k}} Dir(\boldsymbol{\theta}_{n,\pi_n} | \boldsymbol{\alpha}_{n,\pi_n}) d\boldsymbol{\theta}_{n,\pi_n} \\ &= \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\prod_k \Gamma(\alpha_{n,\pi_n,k})} \int \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k} + \alpha_{n,\pi_n,k} - 1} d\boldsymbol{\theta}_{n,\pi_n} \end{aligned}$$

Likelihood:

$$P(\mathcal{D}|\mathcal{G}, \boldsymbol{\theta}) = \prod_n \prod_{\pi_n} \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k}}$$

Recall:

$$Dir(\boldsymbol{\theta}) = \frac{\Gamma(\sum_k \alpha_k)}{\prod_k \Gamma(\alpha_k)} \prod_{k=1}^K \theta_k^{\alpha_k-1}$$

$$\int_{\boldsymbol{\theta} \in \mathcal{S}} \prod_{k=1}^K \theta_k^{\alpha_k-1} d\boldsymbol{\theta} = \left( \frac{\prod_k \Gamma(\alpha_k)}{\Gamma(\sum_k \alpha_k)} \right)$$

Marginal likelihood:

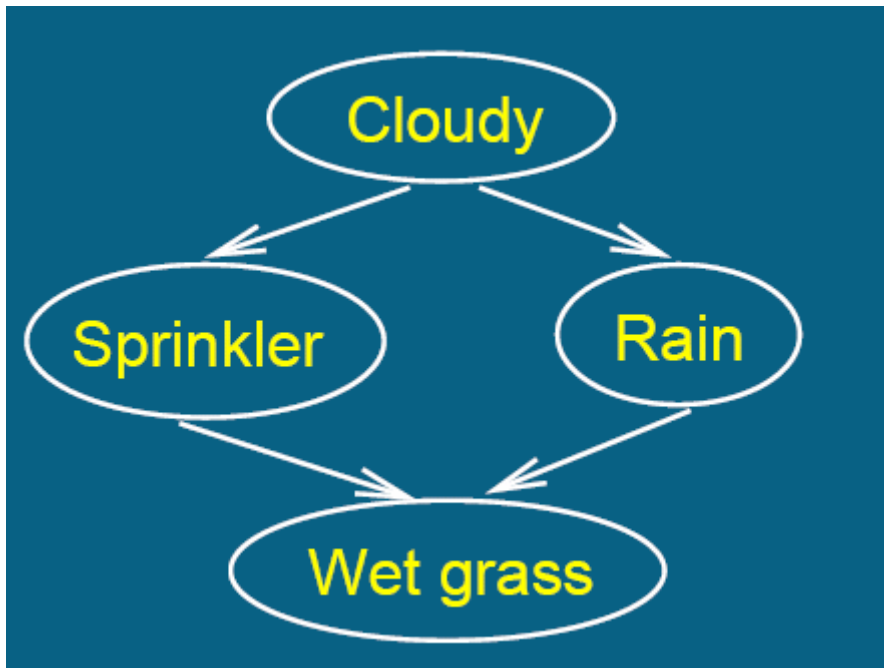
$$\begin{aligned} P(\mathcal{D}|\mathcal{G}) &= \int P(\mathcal{D}|\boldsymbol{\theta}, \mathcal{G}) P(\boldsymbol{\theta}|\mathcal{G}) d\boldsymbol{\theta} \\ &= \prod_n \prod_{\pi_n} \int \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k}} Dir(\boldsymbol{\theta}_{n,\pi_n} | \boldsymbol{\alpha}_{n,\pi_n}) d\boldsymbol{\theta}_{n,\pi_n} \\ &= \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\prod_k \Gamma(\alpha_{n,\pi_n,k})} \int \prod_k [\theta_{n,\pi_n,k}]^{M_{n,\pi_n,k} + \alpha_{n,\pi_n,k} - 1} d\boldsymbol{\theta}_{n,\pi_n} \\ &= \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\prod_k \Gamma(\alpha_{n,\pi_n,k})} \times \frac{\prod_k \Gamma(M_{n,\pi_n,k} + \alpha_{n,\pi_n,k})}{\Gamma(\sum_k [M_{n,\pi_n,k} + \alpha_{n,\pi_n,k}])} \\ &= \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\Gamma(\sum_k [M_{n,\pi_n,k} + \alpha_{n,\pi_n,k}])} \prod_k \frac{\Gamma(M_{n,\pi_n,k} + \alpha_{n,\pi_n,k})}{\Gamma(\alpha_{n,\pi_n,k})} \end{aligned}$$

## BDe score

$$P(\mathcal{D}|\mathcal{G}) = \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\Gamma(\sum_k [M_{n,\pi_n,k} + \alpha_{n,\pi_n,k}])} \prod_k \frac{\Gamma(M_{n,\pi_n,k} + \alpha_{n,\pi_n,k})}{\Gamma(\alpha_{n,\pi_n,k})}$$

## BDe score

$$P(\mathcal{D}|\mathcal{G}) = \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\Gamma(\sum_k [M_{n,\pi_n,k} + \alpha_{n,\pi_n,k}])} \prod_k \frac{\Gamma(M_{n,\pi_n,k} + \alpha_{n,\pi_n,k})}{\Gamma(\alpha_{n,\pi_n,k})}$$



$n$	$\pi_n$	$k$
Clouds	-	on , off
Sprinkler	Clouds = off	on , off
Sprinkler	Clouds = on	on , off
Rain	Clouds = off	on , off
Rain	Clouds = on	on , off
Wet grass	Clouds= off, sprinkler= off	on, off
Wet grass	Clouds= off, sprinkler= on	on, off
Wet grass	Clouds= on, sprinkler= off	on, off
Wet grass	Clouds= on, sprinkler= on	on, off

## BDe score

$$P(\mathcal{D}|\mathcal{G}) = \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\Gamma(\sum_k [M_{n,\pi_n,k} + \alpha_{n,\pi_n,k}])} \prod_k \frac{\Gamma(M_{n,\pi_n,k} + \alpha_{n,\pi_n,k})}{\Gamma(\alpha_{n,\pi_n,k})}$$

How to set the hyperparameters?

## BDe score

$$P(\mathcal{D}|\mathcal{G}) = \prod_n \prod_{\pi_n} \frac{\Gamma(\sum_k \alpha_{n,\pi_n,k})}{\Gamma(\sum_k [M_{n,\pi_n,k} + \alpha_{n,\pi_n,k}])} \prod_k \frac{\Gamma(M_{n,\pi_n,k} + \alpha_{n,\pi_n,k})}{\Gamma(\alpha_{n,\pi_n,k})}$$

How to set the hyperparameters?

Prior network  $\mathcal{G}_0$  and equivalent sample size  $M_0$ . To satisfy event equivalence:

$$\alpha_{n,\pi_n,k} = M_0 P(X_n = k, \pi_n | \mathcal{G}_0)$$

Uninformative case:

$$\alpha_{n,\pi_n,k} = \frac{M_0}{K_n \times \#\pi_n}$$

Number of discretization  
levels for node n

Number of parent  
configurations for node n

BDe: UAI 1994

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Learning Bayesian Networks: The Combination of  
Knowledge and Statistical Data

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BGe: UAI 1995

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Learning Bayesian Networks:  
A Unification for Discrete and Gaussian Domains

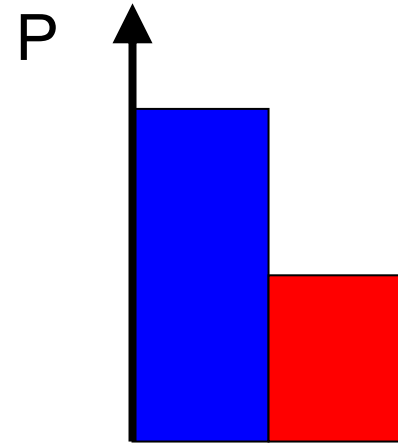
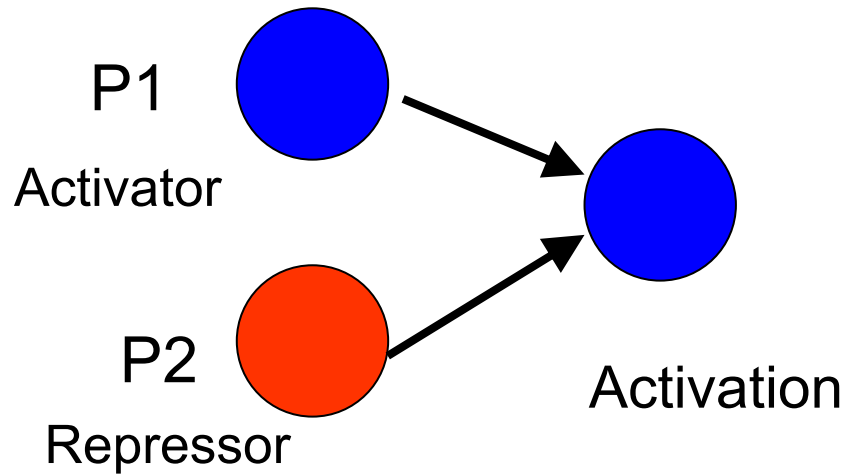
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David Heckerman and Dan Geiger\*

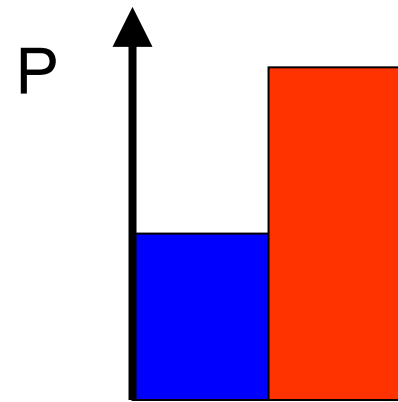
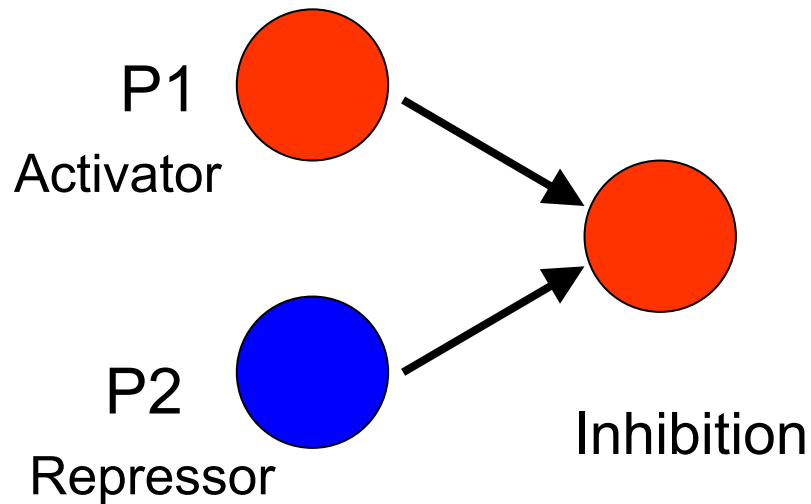
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# BDe: Nonlinear discretized model

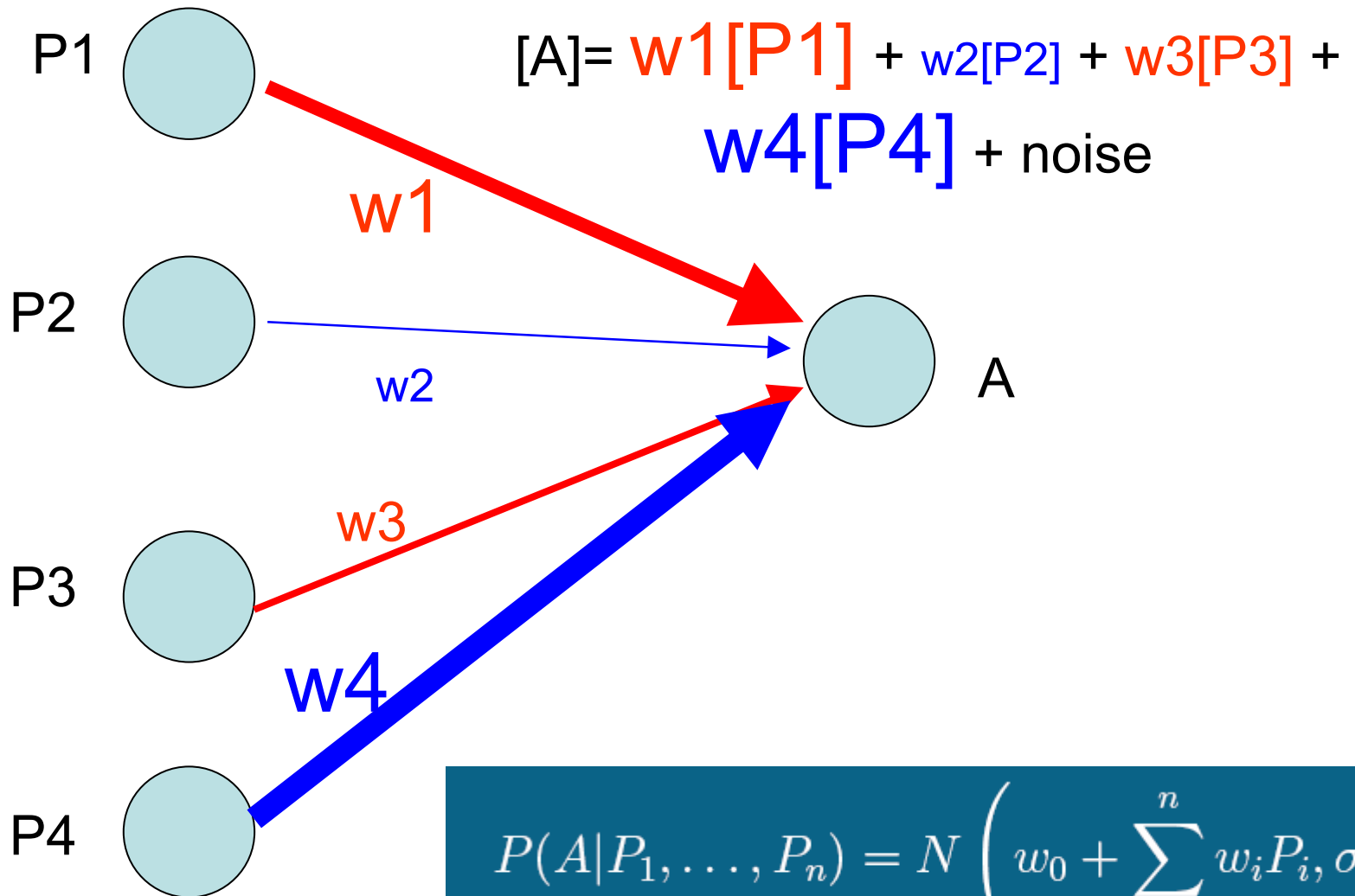


Allow for noise: probabilities



Conditional multinomial  
distribution

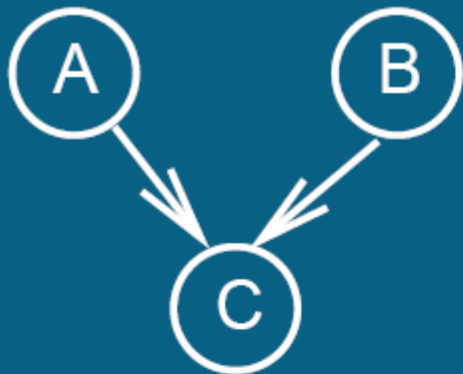
# BGe: Linear model



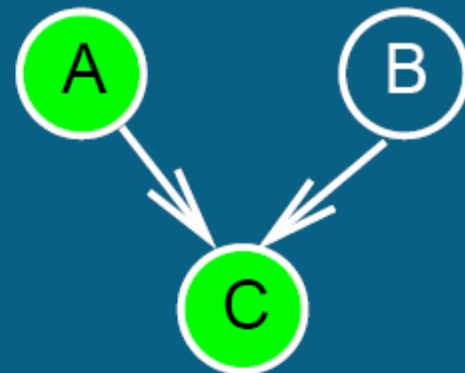
$$P(A|P_1, \dots, P_n) = N \left( w_0 + \sum_{i=1}^n w_i P_i, \sigma^2 \right)$$

## Pros and cons of the two models

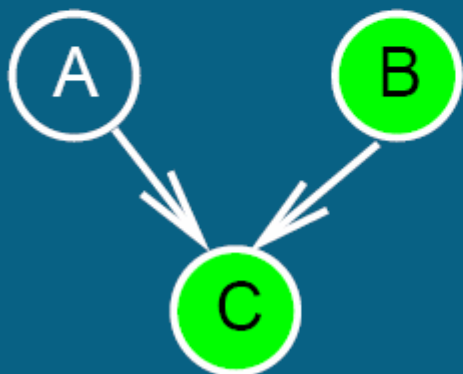
Linear Gaussian model	Multinomial model
<ul style="list-style-type: none"><li data-bbox="110 572 820 729">• Restriction to linear processes</li><li data-bbox="110 872 807 1029">• Original data → no information loss</li></ul>	<ul style="list-style-type: none"><li data-bbox="986 572 1595 629">• Nonlinear model</li><li data-bbox="986 893 1589 1051">• Discretization → information loss</li></ul>



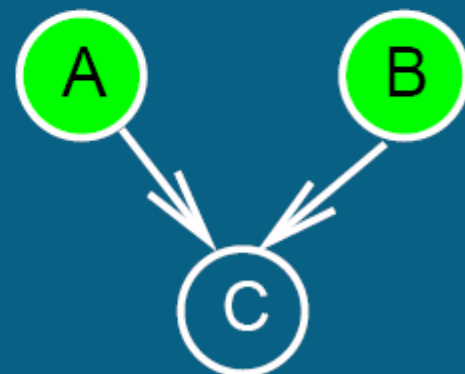
$$C = w_1 A + w_2 B$$



$$C = \uparrow w_1 A + w_2 B$$

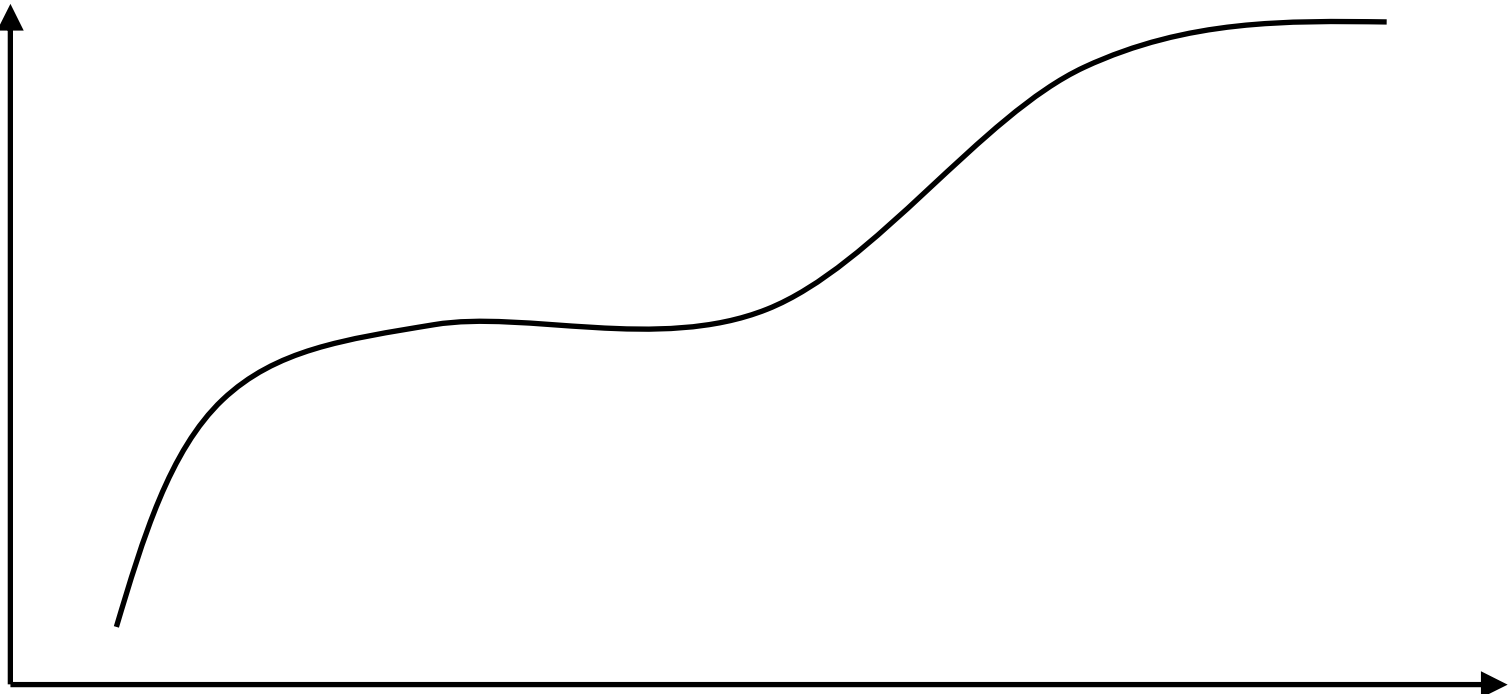


$$C = w_1 A + \uparrow w_2 B$$



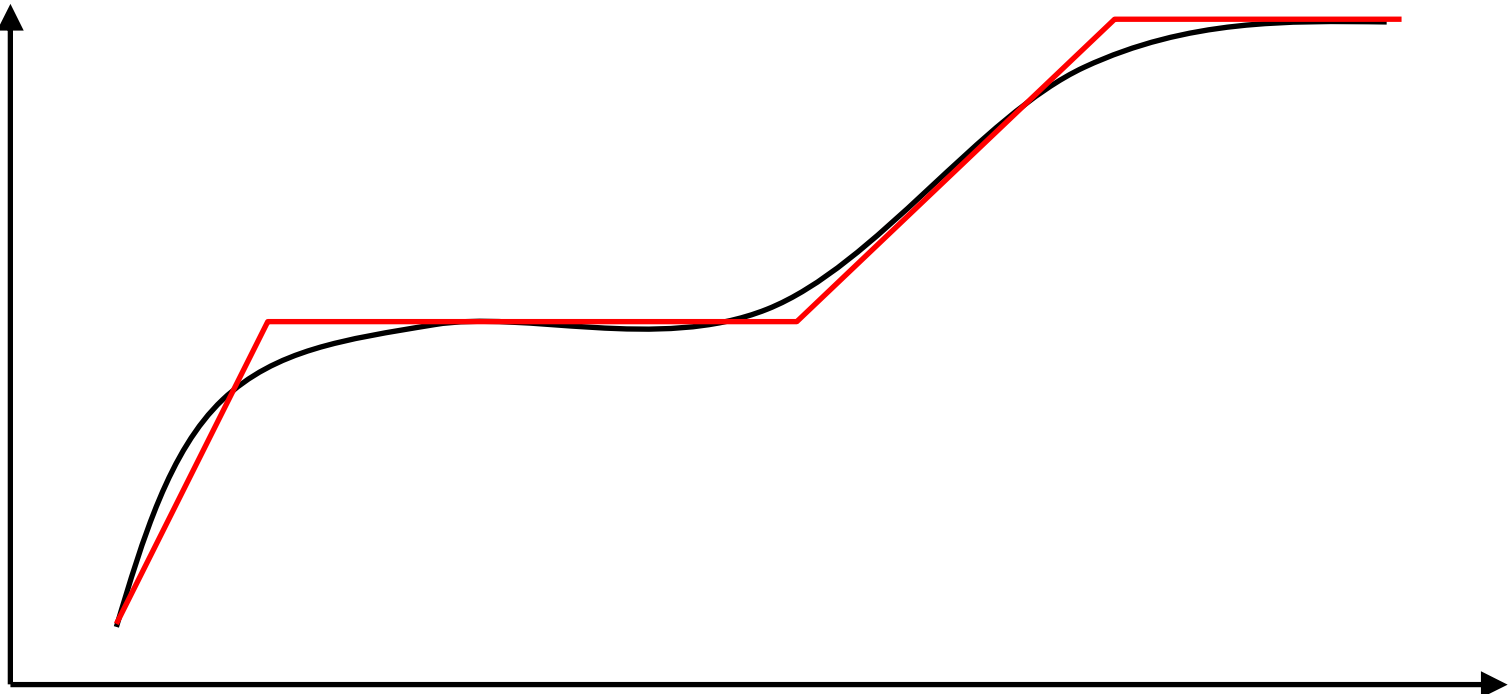
$$C = \uparrow w_1 A + \uparrow w_2 B$$

Can we get an approximate nonlinear model  
without data discretization?



Can we get an approximate nonlinear model  
without data discretization?

Idea: piecewise linear model



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# **Non-stationary continuous dynamic Bayesian networks**

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**NIPS 2009**